DATA CLASSIFICATION BASED ON THE LOCAL INTRINSIC DIMENSION



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Data classification based on the intrinsic dimension

classification and intrinsic dimension

Classification

• hard in high dimension (many variables): computational problems, sampling issues...

- # of independent directions of variation can be lower: intrinsic dimension (ID)
- Accounting for ID can improve classification schemes





Our journey: from classification to intrinsic dimension and back

- We started with a classification problem (clustering algorithm)
- This required accurate ID estimation
- We developed a method to estimate the ID
- We realized than often the ID is not constant within a dataset
- This in turn allows for rough, topologically-based data classification

Density-based clustering

• Reconstruct the probability density of the data



- What is high? Results depend on the chosen density threshold
- Cannot resolve features at different density scales

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Density Peak Clustering

A Rodriguez, A Laio, Science 344, 1492 (2014) M d'Errico, E Facco, A Laio, A Rodriguez, arXiv:1802.10549 (2018)

Cluster around density peaks, i.e. local maxima in the density



Original algorithm: density peaks are far from any point with higher density

Compute for all points min distance from point at higher density $\delta_i = \min_{j:\rho_j > \rho_i} d_{ij}$ Peak are outliers in decision graph ρ_i vs δ_i :

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Density-peak clustering



Points are assigned to peaks by following a path of increasing density leading to one of the peaks.

One jumps from a point to a point with higher density



Density-peak clustering A Rodriguez, A Laio, Science 344, 1492 (2014) M d'Errico, E Facco, A Laio, A Rodriguez, arXiv:1802.10549 (2018)



Points are assigned to clusters by following a path of increasing density leading to one of the peaks.

This assignation rule allows to retrieve clusters of arbitrary shape

K-means

0.5





Short-term coherent pattern detection in fMRI

Density estimation

- Data can be thought of as samples of a density distribution •
- Reconstruct the probability density of the data with proper *density estimator* •
- K-nearest-neighbor: Assume $\rho \approx \text{const}$ in small region around each point •
- For each point *i*, consider its *k* nearest neighbors at • distances $r_{i1}, r_{i2}, r_{i3}, \ldots$
- density= *k*/volume of sphere containing the *k* points •

VI.

$$\rho = \frac{k}{V_{ik}} \qquad \delta\rho = \frac{\sqrt{k}}{V_{ik}} \qquad V_{ik} = \omega_d r_{ik}^d$$

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Density Estimation: PAk

A Rodriguez, M D'Errico, E Facco, A Laio, JCTC (2018)



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$$\rho = \frac{k}{V_{ik}} \qquad \delta \rho = \frac{\sqrt{k}}{V_{ik}} \qquad \qquad V_{ik} = \omega_d r_{ik}^d$$



- Two problems:
- 1) what is right *k*?
- 2) what is right *d*?

Density Estimation: PAk



what is right *k*?

k too small: large error in the estimate

k too large: density is not constant over V

Solution:

- Adapt *k to* each point so that the constant density assumption always holds
- Pointwise Adaptive k-NN (PAk) estimator



Intrinsic dimension



Problem 2): what is right *d*?

The data actually lie on hypersurface of lower dimension than D

The density should be evaluated on this hypersurface



Intrinsic dimension

Example: molecular dynamics





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Intrinsic dimension



Example: images



20x20



 $\mathbb R$

ID estimation: projective approach



Project *D*-dimensional data into lower dimension d: Π^d : $\mathbf{x}_i \in \mathbb{R}^D \mapsto \mathbf{y}_i \in \mathbb{R}^d$

- Try different *d* and evaluate for each a "loss function" $\mathcal{L}(\Pi^d)$
- $\mathcal{L}(\Pi^d)$ measures the "data loss" occurring in the projection. Examples:

 $\mathcal{L}(\Pi^d) = \sum_i ||\mathbf{x}_i - \mathbf{y}_i||^2 \quad \text{preservation of original distance relations}$ $\mathcal{L}(\Pi^d) = \sum_i \mathbf{x}_i \mathbf{x}_i^T - \mathbf{y}_i \mathbf{y}_i^T \quad \text{preservation of original covariance matrix}$

- tradeoff between dimension reduction and data loss
- Problem (1): Computationally burdensome (search for optimal projection for each d)
- Problem (2): robust ID estimates only if $\mathcal{L}(\Pi^d)$ has large gap as a function of d if no gap, the estimation can be rather arbitrary

ID estimation: projective approach

- Example: Principal Component Analysis (PCA)
- Prjoects data onto linear subspace spanned by first *d* eigenvalues of covariance matrix. $X^T X$ Loss: $\mathcal{L}(\Pi^d) = ||\sum \mathbf{x}_i \mathbf{x}_i^T - \mathbf{y}_i \mathbf{y}_i^T||$
- on the villin headpiece simulation:



How can one select an appropriate d

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ID estimation: statistical approach



- assumes that the data are sampled from a distribution with density $\rho(\mathbf{X})$
- distances between points in the dataset follow a scaling law that depends
- on $\rho(\mathbf{X})$ and d
- If the dependence on $\rho({\bf X})$ can be removed, then d can be estimated from the scaling
- Example: correlation dimension
 - The number of points at distance $< \epsilon$ from point *i* scales as $N_i(\epsilon) = \sum_j \theta(d_{ij} < \epsilon) \approx \epsilon^d / \rho(\mathbf{X}_i)$
 - If $\rho(\mathbf{X})$ is constant, $N(\epsilon) = \sum_{ij} \theta(d_{ij} < \epsilon) \sim \epsilon^d / \rho$
 - *d* can be estimated with simple linear fit
- However, when $ho(\mathbf{X})$ is variable the estimation fails dramatically

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ID estimation: TWO-NN

E Facco, M D'Errico, A Rodriguez, A Laio, Scientific Reports 7, 12140. (2017)



- In principle, one should evaluate simultaneously both d and $ho(\mathbf{X})$!
- TWO-NN idea: decouple the estimation problem by finding suitable function of the distances that depends only on d
- Assumption: $\rho(\mathbf{X})$ is constant on the scale of the first two neighbors
- Then if d_{i1}, d_{i2} are distances from 1st and 2nd neighbor of point i,
- their ratio $\mu_i = \frac{d_{i2}}{d_{i1}}$ follows a Pareto distribution: $f(\mu_i) = d\mu_i^{-(d+1)}$
- depends only on *d*, not on $\rho(\mathbf{X})$!
- Collect the μ for each point. Fit their emprical distribution and estimate d
- The ID is inferred from the μ collectively

ID estimation: TWO-NN

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ID estimation: TWO-NN

There are several ways of fitting:

- One can fit the empirical cumulative distribution of μ with $F(\mu) = 1 \mu^{-d}$
- Equivalently, linear fit on $\log(1 F(\mu)) = -d\log\mu$



• If the model is satisfied, then the distribution of the μ_{-} is well fitted (check χ^{2})

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The problem of multiple IDs



If the fit is not good, it means the model fails



- 1) the density is strongly varying even on the scale of the first two neighbors
- 2) the dimension is not uniform in the dataset

The data may lie on several manifolds M_1, \ldots, M_K , each with different ID How to deal with this heterogeneous ID case?

Heterogeneous ID



he data may lie on several manifolds, each with different ID



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- H1) data sampled from manifolds of different ID
- H2) ρ is uniform on scale of the first neighbors
- Under H1), H2) one can still predict the expected distribution of the $~\mu$
- Assume point sampled from M_1, \ldots, M_K with different probabilities $\mathbf{p} = p_1 \ldots p_K$
- mixture of Pareto distributions $P(\mu_i) = \sum_{k=1}^{K} p_k d_k \mu_i^{-d_k-1}$
- The likelihood of the data is $\mathcal{L}(\mu)$

$$\mathcal{L}(\boldsymbol{\mu}|\mathbf{d},\mathbf{p}) = \prod_{i=1}^{N} \sum_{k=1}^{K} p_k d_k \mu_i^{-d_k-1}$$

- Then we can again estimate $\mathbf{d} = d_1 \dots d_K$, $\mathbf{p} = p_1 \dots p_K$
- K is not estimated as a parameter
- Fix K by trying increasing values in $[1, K_{max}]$ and performing a model selection test

- To estimate parameters, fix inferential approach
- A) frequentist: $\mathbf{d}^e, \mathbf{p}^e = argmax(\mathcal{L}(\boldsymbol{\mu}|\mathbf{d},\mathbf{p}))$
- •
- B) Bayesian
 - Fix $P_{prior}(\mathbf{d}, \mathbf{p})$
 - Compute mean $\mathbf{d}^{e}, \mathbf{p}^{e} = \langle \mathbf{d}, \mathbf{p} \rangle_{post}$ $P_{post}(\mathbf{d}, \mathbf{p})$

$$P_{post}(\mathbf{d},\mathbf{p}) \propto \mathcal{L}(\boldsymbol{\mu}|\mathbf{d},\mathbf{p})P_{prior}(\mathbf{d},\mathbf{p})$$

• Because of the sum over k, hard to work with

$$\mathcal{L}(\boldsymbol{\mu}|\mathbf{d},\mathbf{p}) = \prod_{i=1}^{N} \sum_{k=1}^{K} p_k d_k \mu_i^{-d_k-1}$$

- Introduce latent variables $\mathbf{Z} = Z_1, \dots, Z_N$: manifold membership of each point
- Likelihood is seen as marginal over $\mathcal{L}(\boldsymbol{\mu}|\mathbf{d},\mathbf{p},\mathbf{Z}) = \prod_{i=1}^{N} p_{Z_i} d_{Z_i} \mu_i^{-d_{Z_i}-1}$
- Estimate jointly $\mathbf{d}, \mathbf{p}, \mathbf{Z}$
- Heterogeneous ID algorithm (hidalgo)

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Problem: this approach does not work!

Two manifolds of dimension $d_1=4$ and $d_2=5,...,9$ (Gaussian ρ)

estimation of d_1 and d_2 is inaccurate

estimation of Z is completely wrong

Why?

Pareto distributions with different *d* are highly overlapping

The Z assignation is based only on the $\boldsymbol{\mu}$ of each point

Difficult to assing Z if $\boldsymbol{\mu}$ value is not predictive



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We get non-uniform neighborhoods

Neighboring points have different Z



We must assume that the manifolds are separated, with at most a (small) intersection

This implies that the neighborhoods must be approximately uniform

We enforce this through additional term in the likelihood

Let the neighborhood of point *i* be defined by its first *q* neighbors

 n_i^{in} # neighbors with same Z as i n_i^{out} # neighbors with different Z

$$\mathcal{L}(n_i^{in} | \mathbf{Z}) = \frac{\zeta^{n_i^{in}} (1 - \zeta)^{n_i^{out}}}{\mathcal{Z}}$$

 $\zeta > \frac{1}{2}$ Parameter that controls degree of uniformity

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We enforce uniform neighborhoods through **additional term in the likelihood**

$$\mathcal{L}(n^{in}|\mathbf{Z}) = \prod_{i} \frac{\zeta^{n_i^{in}} (1-\zeta)^{n_i^{out}}}{\mathcal{Z}}$$



Now we get correct estimates of both *d*,*p* and *Z*



Hidalgo M Allegra, E Facco, A Laio and A Mira, in prep. (2018)



We achieve a global topological description of the data space

Divide space into regions of uniform intrinsic dimension





Example: molecular dynamics



• for each of the N ~ 32000 configurations, D=32 dihedral angles.

We find four manifolds



The folded state is recognized from its higher ID!

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Example: fMRI time series

- consider ~30000 time series corresponding to BOLD signal of each voxel in an fMRI experiment
- for each of the N ~ 30000 time sieries, D=202 values

We find two manifolds: d=16, d=32



Red: high-ID voxels Blue: "task-relevant" voxels Green: intersection

Task-relevant voxels are in the manifold with higher ID

The low-dimensional manifold nostly includes "noise" voxels

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Example: firms from Compustat

- consider ~8000 firms in the Compustat Database
- for each of the firms, D=31 balance sheet variables

We find four manifolds: d=5, d=6, d=7, d=9

We compute S&P ratings for the different manifolds



Lower dimension tends to have lower ratings!



- The problem of clustering led us to the problem of density estimation; the problem of density estimation led us to the problem of ID estimation
- We developed a reliable ID estimator, TWO-NN, that limits the issue of density variations
- We realized that often the ID is not constant in the dataset: we extended the statistical framework of TWO-NN to comply with this case
- We developed Hidalgo, a method that finds groups of points (manifolds) of different ID in the manifold
- Applications of Hidalgo to real datasets reveal that the topological information given by the ID discriminates points differing in important features

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Thank you for your attention!!